

# Health, Gender and Mobility: Intergenerational Correlations in Longevity over Time

John Parman, University of California - Davis

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# Motivation

- Existing mobility measures have a variety of shortcomings in terms of interpretation and comparison across groups (particularly gender) and over time
- Modern intergenerational income elasticities cannot be estimated for historical time periods and are often sensitive to transitory fluctuations in economic status
- Historical studies have been limited to occupational mobility which limits the samples that can be studied and has ambiguous welfare implications of observed mobility rates
- This paper takes an alternative path to estimate mobility through long term health outcomes addressing many of the limitations of occupational and income mobility studies

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# Brief overview of paper

- Existing mobility measures suggest a decline in American mobility over the past 150 years
- It is still unknown whether this decline existed for females as well as males and how the decline translated into persistence of welfare across generations
- This paper creates a new intergenerational dataset of linked death certificates that allows for estimating intergenerational correlations in longevity
- Longevity is correlated with occupational status both in terms of a child's occupation and father's occupation
- I find strong intergenerational longevity elasticities for both males and females
- These elasticities have been getting stronger for males over the past century but have been stable for females

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# Measures of modern income mobility

- Common approach is to measure intergenerational income elasticities:

$$\ln y_s = \alpha + \rho \ln y_f + \dots$$

- Elasticities typically found to be in the 0.2-0.5 range (see Solon (1999) for survey of estimates)
- Female mobility estimates compare household income of a daughter to household income of parents,  $\rho$  looks slightly smaller for females

# Measures of modern income mobility

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Drawbacks of intergenerational income elasticities:

- Sensitive to transitory fluctuations in income (this can have big effects on estimates, see Mazumder (2005))
- Reliance on household income as a measure of individual welfare
- Lack of sufficient income data to estimate historical trends

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# Measures of historical occupational mobility

- Historical income data is rare
- Studies have relied on linking individuals across censuses and measuring occupational transitions
- Results suggest that US exhibited high levels of occupational mobility in the 19th century that have fallen over time (Ferrie (2005), Ferrie and Long (2006, 2009))

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# Measures of historical occupational mobility

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Drawbacks of intergenerational occupational mobility:

- Difficult to translate occupational status to well being
- Major changes in occupational structure and nature of occupations over time
- Occupational change over career
- Difficult to interpret historically for females
- Strategy can't be used for females due to changing surnames

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# Motivating health as a measure of mobility

- Annual income and occupation are imperfect measures of long term economic status
- Both suffer from the effects of intragenerational mobility
- Both have issues of how to treat the spouse's income and occupation
- Occupation in particular is difficult to map into overall welfare

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# Motivating health as a measure of mobility

- A greater concern for our purposes is the inability of these measures to address female mobility
- Income mobility based on household income is a tricky measure of female welfare if the distribution of household resources (and work) is changing over time
- Occupational mobility based on female occupation has clear problems if looking at historical mobility patterns
- Occupational mobility based on spouse occupation is limited by surname changes

# Why should we care about separate female mobility estimates?

- The obvious: females are half the population
- More importantly, there are a variety of reasons we would expect historical mobility patterns of females to differ from the historical mobility patterns of males
- The forces shaping the economic status of women and men have followed very different paths historically

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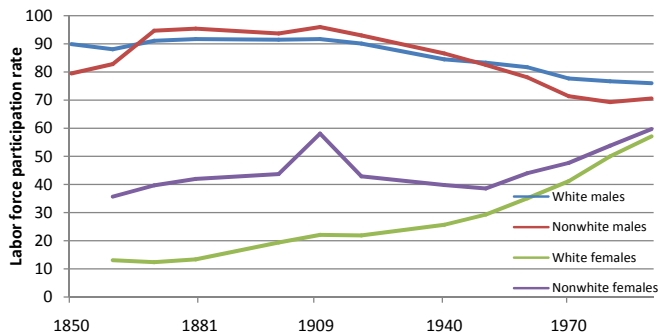
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# Why should we care about separate female mobility estimates?



Labor force participation rates by gender and race, 1850-1990. Source: Sobek (2001).

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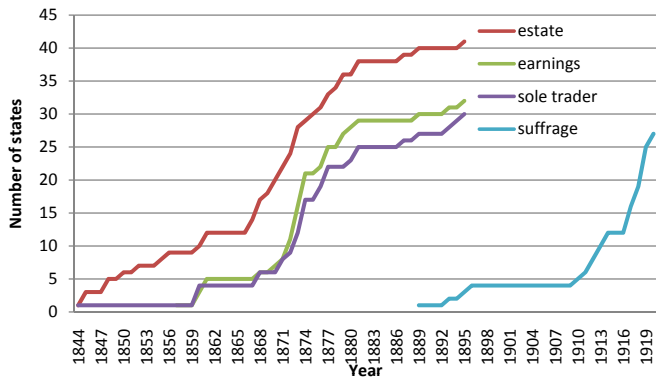
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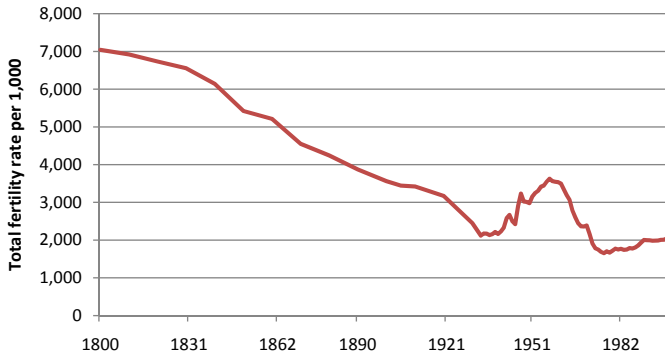
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# Why should we care about separate female mobility estimates?



Number of states with women's property rights and suffrage laws, 1844-1920. Sources: Khan (1996), Rusk (2001).

# Why should we care about separate female mobility estimates?



Total fertility rate per 1,000 for white females, 1800-1998.  
Source: Historical Statistics of the United States.

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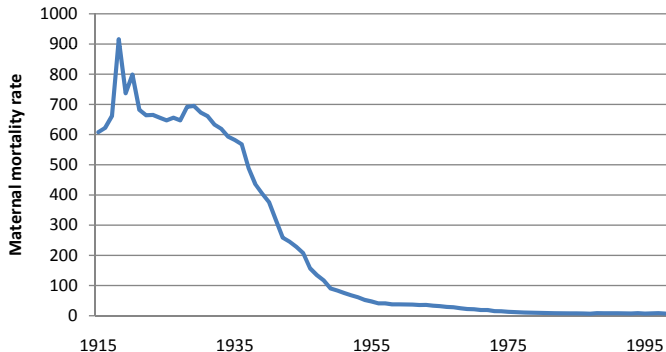
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# Why should we care about separate female mobility estimates?



Maternal mortality rate per 100,000 live births, 1915-1998.  
Source: US Public Health Service vital statistics.

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# Health as a measure of intergenerational mobility

The theoretical appeal of long term health outcomes as a measure of mobility:

- Long term health is a direct measure of quality of life, not a measure of a means to a particular quality of life
- Health outcomes have a consistent metric that allows for meaningful comparisons over time and across gender
- Health outcomes are individual outcomes (we don't need to understand how the allocation of household resources works)
- Lexicographic nature of health and economic success: health is a first order concern

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# Health as a measure of intergenerational mobility

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The practical appeal of long term health outcomes as a measure of mobility:

- Historical evidence on certain health outcomes is more prevalent than income and wealth data
- Annual data rather than decennial data
- Less affected by transitory fluctuations over life cycle
- Reported for males and females
- Better reporting of parental information

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# Death certificates as a data source

- The basic approach will be to match the death certificates of parents and children
- Death certificates have several useful features:
  - Cover entire state populations
  - Extend back to the 19th century for several states
  - Contain excellent long term health data and occupation information
  - More detailed information for linking than the census provides
- Focus for initial study will be the death certificates for Mecklenburg County, NC
  - North Carolina has detailed death records available from 1909 to 1975
  - Mecklenburg has significant urban and rural populations
  - Initial sample is a 10% sample of deaths in the county

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# North Carolina death certificate, 1910

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WRITE PLAINLY, WITH UNFADING INK--THIS IS A PERMANENT RECORD.

N. B.--This form of information should be carefully reported. AGES should be stated EXACTLY. PHYSICIANS should state ONSE OF DISEASE, and the cause of death, in full, and the name of the physician. The "Special Information" for persons dying away from home should be given in full in every instance.

Registered No. 21

STATE OF NORTH CAROLINA.

STATE BOARD OF HEALTH--DIVISION OF VITAL STATISTICS.

PLACE OF DEATH. 74

## CERTIFICATE AND RECORD OF DEATH.

Town of Greenville, N.C.

(No. \_\_\_\_\_ Street.)

(If death occurred in a hospital or institution give its NAME instead of street and number.)

(If death occurs away from usual residence give full address for "Special Information.")

FULL NAME Ellie Luora Fleming

### PERSONAL AND STATISTICAL PARTICULARS.

SEX <u>Female</u>	COLOR <u>White</u>
DATE OF BIRTH <u>Jan 13 1874</u> (Month) (Day) (Year)	
AGE <u>35</u> years <u>11</u> months <u>21</u> days.	
SINGLE, MARRIED, WIDOWED, OR DIVORCED <u>Married</u>	
BIRTHPLACE (State or country) <u>Pitt County</u>	
NAME OF FATHER <u>Francis M. Smith</u>	
BIRTHPLACE OF FATHER (State or country) <u>Pitt County</u>	
MARIED NAME OF MOTHER <u>Amanda Watston</u>	
BIRTHPLACE OF MOTHER (State or country) <u>Pitt County</u>	
OCCUPATION <u>none</u>	

THE ABOVE STATED PERSONAL PARTICULARS ARE TRUE TO THE BEST OF MY KNOWLEDGE I BELIEVE.

(Informant) J. T. Smith  
Greenville, N.C.

Filed 1/6 1910 Jan 6 3pm

### MEDICAL CERTIFICATE OF DEATH.

DATE OF DEATH  
Jan 5 1910  
(Month) (Day) (Year)

I HEREBY CERTIFY, That I attended deceased from

Nov 22 1899 to Jan 5, 1910.

That I last saw her alive on Jan 5, 1910.

and that death occurred, on the date stated above, at 11:30 P. M.

THE CAUSE OF DEATH was as follows:

Acute Parenchymatous Nephritis -

(duration) 43 days.

Contributory Pregnancy

(duration) \_\_\_\_\_ days.

(Signed) Jan E. Forbes M. D.  
1-6-1910 (Address) Greenville, N.C.

SPECIAL INFORMATION only for Hospitals, Institutions, Transients, or Boarding Houses.

Place or Usual Residence \_\_\_\_\_ How long at \_\_\_\_\_ Place of death? \_\_\_\_\_ days.

Where was disease contracted, if not at place of death? \_\_\_\_\_

PLACE OF BURIAL OR REMOVAL  
Cherry Hill Cemetery

DATE OF BURIAL  
1/6 1910

UNDERSTAKER  
John Howard and Company, Greenville, N.C.

# North Carolina death certificate, 1975

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FEB 7 1975  
DISTRICT NO. 05095 SOCIAL NO. 2

NORTH CAROLINA STATE BOARD OF HEALTH  
OFFICE OF VITAL STATISTICS  
CERTIFICATE OF DEATH 2525

NAME OF DECEASED: Irvin Nathan CGK  
DATE OF BIRTH: January 1, 1975  
SEX: Male COLOR: White STATE OF BIRTH: North Carolina DATE OF DEATH: Sept. 24, 1988  
AGE IN YEARS: 87

PLACE OF DEATH: Mecklenburg County, Charlotte, North Carolina  
USUAL RESIDENCE: Mecklenburg County, Mecklenburg  
HOSPITAL OF INSTITUTION: Memorial Hospital, Charlotte

MARRIED: Married  
SPOUSE: Connie Bell Auman  
STREET ADDRESS: 140 Hunter Lane  
CITY: Charlotte

FATHER'S NAME: Clarkson J. Cox MOTHER'S MARRIED NAME: Sophrona Spencer  
DECEASED'S HOME ADDRESS: Mrs. Connie Bell Auman, 140 Hunter Lane, Charlotte, North Carolina 28211

CAUSE OF DEATH: Myocardial infarction  
INTERVALS: 12-14 hours  
ARTERIO-SCLEROSIS: 4 and 5 years

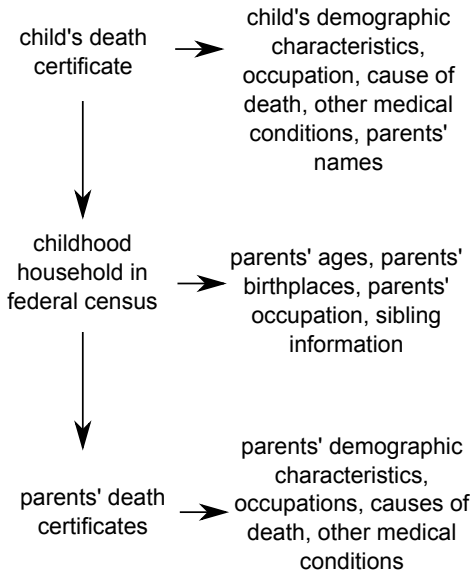
OTHER SIGNIFICANT CONDITIONS: Coronary atherosclerosis, Septicemia  
ACCUENT, SUICIDE, HOMICIDE, OR UNDESIGNED: None

CERTIFICATION - PHYSICIAN: 6/4 - 69  
CERTIFICATION - MEDICAL EXAMINER OR ACTING MEDICAL EXAMINER: WILLIAM C. SUGG, JR., M.D.  
1900 BRUNSWICK AVENUE, CHARLOTTE, N. C. 28207

BURIAL: Burial, 1/3/75, Sharon Memorial Park, Charlotte, North Carolina

DATE OF LOCAL REGISTRATION: 8/19/75

# Linking children to their parents



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# Linking success rates

	All individuals	Males	Females
Number of individuals in initial sample	12,317	6,849	5,468
Number of individuals born before 1930	10,104	5,489	4,615
<u>Linking child to federal census</u>			
% of individuals born before 1930 not found in census	61.3%	60.8%	63.0%
% found but not living with parents	0.2	0.1	0.2
% found living with mother but not father	0.6	0.5	0.2
% found living with father but not mother	0.3	--	0.1
% found living with both parents	37.6	38.5	36.5
<u>Linking parents to death certificates</u>			
% of individuals found in census linked to father's death certificate only	18.6%	20%	17.2%
% linked to mother's death certificate only	17.9	18.7	17.4
% linked to both parents' death certificates	20.1	20.2	17.6
Number of individuals matched to at least one parent's death certificate	1521	872	649

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# Characteristics by linking outcome

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	<u>Sons</u>		<u>Daughters</u>	
	Matched to at			
	Not matched to either parent	least one parent	Not matched to either parent	least one parent
Mean life span	58.6	60.9	61.9	64.7
Standard deviation of life span	(20.0)	(15.8)	(21.5)	(17.3)
% born in North Carolina	56%	83%	56%	80%
% born in South Carolina	23	8	23	12
% in farming	15	17	2	0.4
% in skilled/semi-skilled occupations	23	26	26	30
% in textiles	14	15	8	11
% in unskilled occupations	21	8	52	33
% in white collar occupations	28	35	13	26

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# Measures of health from death certificates

The standard information on the death certificates offers several ways to measure health:

- Date of death and date of birth allow for direct calculation of life span
- Interval between onset and death for medical conditions allows for calculating healthy life span (or years in poor health)
- Cause of death and other significant conditions allow for disease-specific measures:
  - Disease incidence
  - Age at incidence
  - Duration of disease
- All of these measures can be calculated for children and their parents

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# Measures of health from death certificates

The initial results will focus on longevity. While this is the simplest measure to begin with, even longevity has a few complicating factors:

- There are two different ways to calculate longevity: using the date of birth given on the death certificate or using the age given on the federal census
- In a perfect world, these would give the same result but in practice they are often different
- This suggests a measurement error problem that will bias any intergenerational elasticities toward zero
- If one birth year is better than another, we should discard the poor one
- If both are equally likely to be misreported, we should average the two

# Age misreporting by gender and race

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	Death certificate birth year - census birth year			Absolute value of (death certificate birth year - census birthyear)		
	Child	Mother	Father	Child	Mother	Father
All	-.29 (1.58)	-.58 (2.45)	-.44 (2.45)	.98 (1.27)	1.61 (1.93)	1.55 (1.95)
Male	-.39 (1.47)	-.57 (2.46)	-.53 (2.42)	.96 (1.17)	1.66 (1.89)	1.54 (1.94)
Female	-.21 (1.62)	-.59 (2.48)	-.38 (2.43)	.98 (1.31)	1.59 (1.99)	1.53 (1.93)
White	-.47 (1.20)	-.79 (2.05)	-.57 (2.20)	.83 (.99)	1.41 (1.68)	1.42 (1.77)
Black	.56 (2.55)	1.08 (4.21)	.48 (3.85)	1.73 (1.95)	3.39 (2.71)	2.68 (2.94)

Notes: Census birth year is calculated by subtracting the age reported in the census from the year of the census. This means that the imputed birth year may be one year off of the death certificate birth year simply because the individual has not reached their birthday by the time of the census.

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# Age heaping by gender and race

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Percentage of observations with a last digit of 0 or 5

	Child's census age	Father's census age	Mother's census age	Child's death certificate birth year	Father's death certificate birth year	Mother's death certificate birth year
All	18.7%	24.4%	23.7%	20.4%	18.6%	19.7%
Male	17.8	24.3	23.0	19.3	17.3	20.6
Female	19.9	24.5	24.6	18.5	20.6	18.6
White	18.9	23.4	22.9	19.5	18.7	19.7
Black	17.6	29.3	27.8	17.9	18.5	20.4

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# Male occupational distribution by cohort

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## Distribution of occupations within cohort

Decade of birth	Skilled, semi-			
	Farmer	skilled	Unskilled	White collar
1830s	67%	17%	0%	17%
1840s	75	16	4	5
1850s	68	11	3	19
1860s	69	11	8	12
1870s	57	16	8	18
1880s	44	22	10	24
1890s	30	27	11	32
1900s	19	24	11	45
1910s	10	45	10	34
1920s	5	35	10	50

Note: Distributions are based on all males in the sample including both fathers and male children.

# Longevity and occupational status

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Occupational category	Daughter's lifespan		
	Son's lifespan by son's occupation	Son's lifespan by father's occupation	by father's occupation
Farmer	68.2 (15.0)	64.7 (13.9)	67.6 (16.6)
Skill, semi-skilled	62.0 (13.5)	56.8 (16.9)	64.1 (19.2)
Unskilled	51.7 (17.8)	53.2 (18.1)	55.4 (17.5)
White collar	60.4 (12.2)	58.4 (15.3)	61.5 (16.4)
Number of observations	360	414	316

Notes: Standard deviations are given in parentheses. Lifespan is calculated using the year of death from the death certificate and the year of birth implied by the age given on the federal census.

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# Longevity and occupational status

	Sons	Daughters
Child's birth year	26.157*** (8.842)	28.860*** (10.879)
(Child's birth year) <sup>2</sup>	-0.007*** (0.002)	-0.008*** (0.003)
<u>Son's occupation dummies</u>		
Skilled/semi-skilled	2.818** (1.326)	--
White collar	1.026 (1.243)	--
Unskilled	-8.419*** (2.230)	--
<u>Father's occupation dummies</u>		
Skilled/semi-skilled	-2.260 (1.508)	-2.721 (1.985)
White collar	-2.291 (1.561)	-1.339 (1.793)
Unskilled	-4.494** (2.140)	-4.426* (2.502)
Constant	-24,111.147*** (8,398.513)	-26,560.912** (10,330.568)
Observations	545	411
R-squared	0.44	0.47

Standard errors in parentheses. \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%. Omitted occupational dummy is farmer for both son and father.

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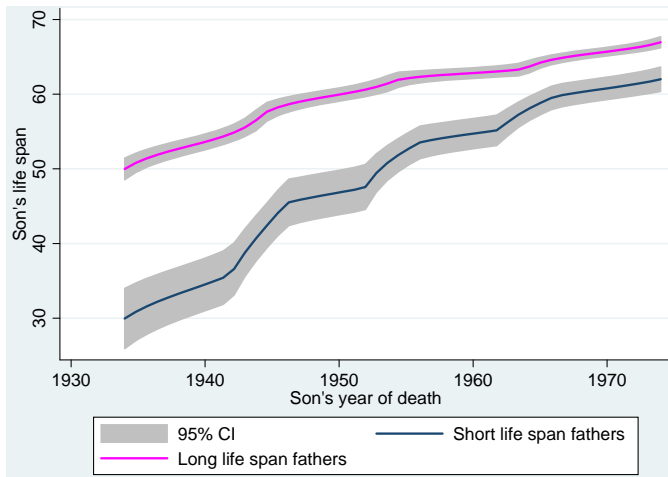
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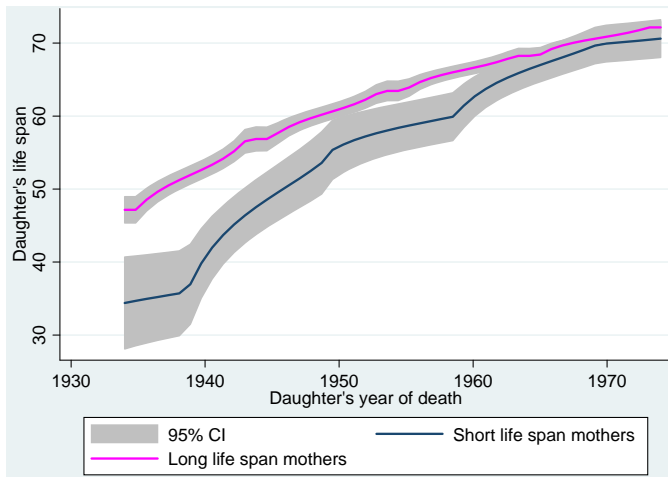
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# Estimating mobility with longevity

- Longevity offers us a nice, continuous variable that measures welfare
- To measure mobility with longevity, the same approaches used for modern income mobility studies can be applied
- The basic idea is to estimate an intergenerational life span elasticity:

$$\ln L_{i,c} = \alpha + \rho \ln L_{i,p} + X_i \beta + \varepsilon_i$$

- $L_{i,c}$  and  $L_{i,p}$  are the life spans of the child and parent for observation  $i$ , respectively
- $X_i$  includes a polynomial in the child's birth year and a polynomial in parent's birth year
- Sons and daughters are treated as separate samples

# Intergenerational longevity elasticities for males

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**Estimates of intergenerational health elasticities for sons and their fathers, log of son's life span as the dependent variable.**

	(1)	(2)	(3)
ln(father's life span)	0.283*** (0.063)	0.274*** (0.066)	0.214*** (0.075)
Son's birth year	0.995*** (0.183)	0.692*** (0.245)	0.634** (0.247)
(Son's birth year)^2	-0.000*** (0.000)	-0.000*** (0.000)	-0.000*** (0.000)
Father's birth year		0.362* (0.190)	0.257 (0.200)
(Father's birth year)^2		-0.000* (0.000)	-0.000 (0.000)
ln(Father's life span) x (son's birth year-1900)/10			0.093* (0.056)
Constant	-930.676*** (174.020)	-980.706*** (175.897)	-789.322*** (210.251)
Observations	586	586	586
R-squared	0.38	0.38	0.38

Notes: Standard errors in parentheses. \* significant at 10%, \*\* significant at 5%, \*\*\* significant at 1%.

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# Intergenerational longevity elasticities for females

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**Estimates of intergenerational health elasticities for daughters and their mothers, log of daughter's life span as the dependent variable.**

	(1)	(2)	(3)
ln(mother's life span)	0.185*** (0.061)	0.203*** (0.064)	0.210*** (0.068)
Daughter's birth year	0.795*** (0.176)	0.760*** (0.274)	0.772*** (0.277)
(Daughter's birth year) <sup>2</sup>	-0.000*** (0.000)	-0.000*** (0.000)	-0.000*** (0.000)
Mother's birth year		0.053 (0.246)	0.080 (0.258)
(Mother's birth year) <sup>2</sup>		-0.000 (0.000)	-0.000 (0.000)
ln(mother's life span) x (daughter's birth year-1900)/10			-0.019 (0.055)
Constant	-737.596*** (167.016)	-754.574*** (170.256)	-799.105*** (215.053)
Observations	425	425	425
R-squared	0.41	0.41	0.41

Notes: Standard errors in parentheses. \* significant at 10%, \*\* significant at 5%, \*\*\* significant at 1%.

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# Intergenerational longevity elasticities using both parents' life spans

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## Life span regressions including both parent's life spans, log of child's life span as dependent variable.

	Sons	Daughters
ln(father's life span)	0.359*** (0.082)	0.090 (0.095)
ln(mother's life span)	0.157** (0.076)	0.320*** (0.086)
Child's birth year	1.516*** (0.279)	0.805** (0.362)
(Child's birth year)^2	-0.000*** (0.000)	-0.000** (0.000)
Constant	-1,427.452*** (265.344)	-749.339** (343.910)
Observations	293	215
R-squared	0.41	0.41

Notes: Standard errors in parentheses. \* significant at 10%, \*\* significant at 5%, \*\*\* significant at 1%.

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## Extensions: Coverage of sample

- Immediate task is to expand the sample coverage: more North Carolina counties, include South Carolina data, more observations at all years
- The larger sample will allow for looking at urban-rural differences and racial differences in mobility patterns
- Better coverage over time will allow for a better treatment of time trends in mobility (especially with regard to women's rights advances)
- If willing to focus only on health (not occupation) several other states can be added to the sample

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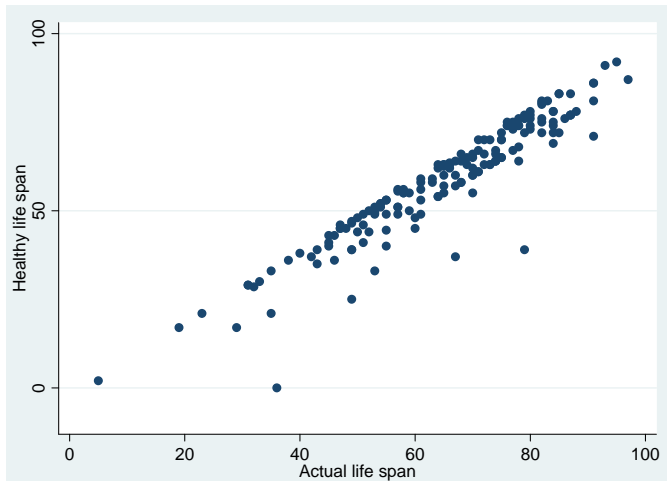
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# Extensions: Healthy life span and years in poor health



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## Extensions: Linking to adult household

- The sons and daughters in our sample can be found as adults in the federal census searching by name, year of birth, state of birth and spouse's name
- Once linked to the federal census, we can add a variety of new variables:
  - Age at marriage
  - Age at first child
  - Number of children
  - Occupation and spouse's occupation
- This would allow for constructing occupational mobility measures for women using spouse's occupation and father's occupation
- Mobility rates could be explained as a function of marriage patterns, fertility patterns, and labor market decisions